

*Report on*

***Exploring Geographic Scale for  
Cancer Incidence Mapping***

*Demonstration Project on*

***Geographic Patterns of Cancer Incidence and  
Environmental Factors in New Jersey***

*Program 03074, Environmental and Health Effects Tracking  
National Center for Environmental Health  
Centers for Disease Control and Prevention*

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**In cooperation with the  
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## **Overview of the Demonstration Project**

The New Jersey Department of Health and Senior Services (NJDHSS) was awarded funding from the Centers for Disease Control and Prevention (CDC) to conduct three demonstration projects under the program, “Environmental and Health Effects Tracking,” in cooperation with the New Jersey Department of Environmental Protection (NJDEP). The purpose of these demonstration projects is to develop and evaluate methods for linking ongoing, existing health effects and human exposure surveillance systems with existing systems for monitoring environmental hazards and exposures.

One of the three demonstration projects by NJDHSS and NJDEP is to link cancer incidence data with data on environmental hazards and exposures. Environmental factors are known or suspected to play an important role in the etiology of several cancer types. This demonstration project will allow NJDHSS and NJDEP to proactively evaluate the geographic relationships among the incidences of selected cancer types and specific environmental hazards or exposures.

The project was conducted in two phases. Phase 1 involves identification of specific cancer types of interest, and descriptive analysis of incidence data for these cancers, specifically for temporal trends and spatial patterns. The second phase involves the linkage of the cancer incidence and environmental databases to examine specific relationships suggested in the first phase. Separate reports describe the methods, findings and conclusions of these phases of the demonstration project.

During Phase 1, the interagency study team began consideration of issues related to geographic scale in the mapping of cancer incidence data. This report discusses the findings of that effort.

This demonstration project was conducted by the Environmental Public Health Tracking Project (EPHT) in Consumer and Environmental Health Services, NJDHSS, in partnership with Cancer Epidemiology Services (CES) and the New Jersey Department of Environmental Protection (NJDEP).

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## Summary

This report explores the presentation and mapping of actual disease rates or relative rate information at small geographic scales. Specifically this report considers the question of what should be an appropriate geographic incidence scale for presentation and mapping of cancer incidence relative rate data in New Jersey, given that cancer types occur with widely varying frequency, and given the large variation in population size distributions of population units in the state.

The report sequentially discusses three fundamental questions relevant to the geographical presentation and mapping of cancer incidence:

- *What is a practical minimum limit on the expected number of cancer cases in a mapped geographic unit that would generate reliable rate ratios?*
- *What size population (or person-time) is needed to generate a practical minimum number of cases for mapping, at varying cancer rates?*
- *How does the distribution of real population units in New Jersey compare to these population sizes?*

The report concludes that in New Jersey, most cancer types could be practically mapped for all counties and for a small number of New Jersey municipalities using 5 year incidence intervals. For most municipalities and some census tracts, reliable rates could only be mapped for the most common cancers (prostate, breast, lung and bronchus, and colon and rectum) using 5 year incidence intervals. For the many cancer groupings with annual incidence rates between 10 and 50 newly diagnosed cases per 100,000 population, the report concludes that five year incidence rates or relative rates could only be reliably mapped at the county level, and for perhaps the 25 or 30 largest municipalities in the state.

## **Introduction**

In 2003, the New Jersey Department of Health and Senior Services (NJDHSS) convened a Cancer Cluster Task Force to examine trends in cancer incidence and mortality, evaluate cancer cluster investigation protocols, and make recommendations for implementation of best practices.

Task Force members, representing government, universities, and public health advocates, expressed interest in the mapping of incidence data at a smaller geographic scale (e.g., municipality or census tract) than typically presented (county or large municipality). One of the reasons for presentation of data at smaller geographic scales is to examine it for local clustering. Mapping of data might allow for the detection of locally elevated rates or other geographic patterns that might point to undiscovered environmental or socioeconomic risk factors. In part, the interest in small scale disease data is also motivated by a community's "right to know" local disease rates.

Also in 2003, the NJDHSS began conducting demonstration projects under the Environmental and Health Effects Tracking (EPHT) project cooperative agreement funded by the U.S. Centers for Disease Control and Prevention (CDC). The Cancer Team established under the New Jersey EPHT project decided to explore the question of appropriate mapping scale for cancer. One of the goals of EPHT is to make public health and environmental data more widely available to the public and researchers, which coincides with the interest expressed by the members of the Cancer Cluster Task Force.

However, there is concern in the public health community about the dissemination and display of information on local disease counts or rates (Rudolph et al., 2006). There are two reasons for this concern: 1) potential violation of ethical and legal protections of the confidentiality of personal health information; and 2) statistical unreliability of the information. Several approaches to decreasing the disclosure risk of confidential data have been proposed and are in use, including administrative controls on access to data, and aggregation of data with rules for cell suppression. Similarly, there are different approaches to increasing statistical reliability of data, including aggregation and geographic smoothing techniques (Rudolph et al, 2006). This report considers the question of statistical reliability only, specifically in relation to the handling of count, rate or relative rate information based on small numbers of cancer cases.

Small area analyses may produce a high proportion of rates or rate ratios that are elevated by chance alone (Neutra et al., 1992; Weinstein and Klotz, 2000). Mapping of unreliable rates may mislead by falsely highlighting areas by chance, or masking an important underlying pattern (Pickle, 2000). While geographic smoothing techniques may be very useful for the visual impact of dampening random variation and revealing geographic patterns, these techniques may not satisfy the "right to know" local disease rates.

This report explores the presentation and mapping of actual disease rate or relative rate information at small geographic scales. Specifically this report considers the

question of what should be an appropriate geographic incidence scale for presentation and mapping of cancer incidence relative rate data in New Jersey, given that cancer types occur with widely varying frequency, and given the large variation in population size distributions of population units in the state.

In practice, public health agencies have used a variety of rules for inclusion of statistics based on low case counts. For example, in its cancer mortality maps and tables, the National Cancer Institute considers data to be too sparse for display if there are fewer than 6 observed deaths, fewer than 12 observed deaths if the rate is not statistically different from the U.S., or fewer than 6 expected deaths (NCI, 2006). The NCI's SEER program requires at least 16 observed or expected cases (Bleyer et al., 2006). In its county-level cancer incidence maps and tables, the New York State Department of Health marks as unstable any rate that is based on fewer than 20 observed cases (NYSDH, 2006).

## Approach

An initial question to consider is:

- *What is a practical minimum limit on the expected number of cancer cases in a mapped geographic unit that would generate reliable rate ratios?*

One approach is to identify a minimum limit such that it would be rare for some predesignated multiple (e.g., a 50% increase, or a doubling) of the expected rate to be exceeded in a geographic unit by chance alone. Based on the Poisson distribution, we examined the cumulative probabilities of observed numbers for a series of expected numbers ranging from 1 to 20 (Figure 1 and Appendix Table 1). For all possible observed numbers for each expected number, we then computed ratios of observed to expected numbers (Figure 2 and Appendix Table 2). This permitted us to determine the expected number that would result in a sufficiently small probability (e.g., 0.05, or 0.001) of an observed number exceeding a rate ratio (RR) of a pre-determined level (e.g.,  $RR \geq 1.5$  or  $\geq 2.0$ ). These probabilities could be interpreted as the proportion of "false positives" in tables or thematic mapping. This proportion, of course, should be minimized to the degree practical.

Once this question has been answered, a second question to consider is:

- *What size population (or person-time) is needed to generate a practical minimum number of cases for mapping, at varying cancer rates?*

Finally:

- *How does the distribution of real population units in New Jersey compare to these population sizes?*

## **Practical Minimum Limit on Number of Cases**

Table 1 shows the probability of a mapped unit having a RR above 1.5 or 2.0 by chance alone, for a series of expected numbers from 2 to 20 (only even numbers shown). Suppose you wanted to produce map in which there is a less than 5% probability that mapped unit RR values exceed 1.5 by chance. From Table 1, you can see that a RR of 1.5 would be exceeded in less than 5% of observations by chance alone (a 5% false positive rate), when expected numbers are *10 or more*. An expected number of 10 would result in less than 0.2% of observations exceeding a RR of 2.0 by chance.

Suppose that you wanted a map in which there was less than a 1% chance of mapped unit RR values exceeding 1.5 by chance. In this case you would need an expected number that is larger than 20. Similarly, suppose you wanted to produce a map in which there is a less than 1% probability that mapped unit RR values exceed 2.0 by chance. From Table 1, you can see that a RR of 2.0 would be exceeded in less than 1% of observations by chance alone when the expected number is *6 or more*. However, an expected number of 6 would result in about 8% of observations exceeding a RR of 1.5 by chance.

A practical minimum limit of 10 or more expected cases may be chosen to ensure low probabilities of “false positives” and reliability for presentation of small area rates. A larger minimum number, perhaps an expected number of 16, would result in fewer “false positives.” Minimum expected numbers larger than about 16 do not seem to appreciably reduce the false positive proportion. This suggests that selecting a minimum expected number in the range of 10 to 16 cases is a reasonable choice.

## **Population Sizes to Generate Minimum Numbers and Comparison to New Jersey Population Units**

For surveillance of incidence data for temporal-geographic clustering of cancers, we assumed that a five year interval would be a reasonable time scale for analysis of data. Tables 2 and 3 show the expected number of cases generated over a five-year interval for a range of incidence rates and for a range of population sizes. Table 3 also shows example cancer types within incidence rate ranges for reference. For the many cancer types whose incidence rates exceed about 10 per 100,000 per year, reliable rates based on expected numbers in the 10 to 16 range can be mapped for population units exceeding 50,000 total persons (or 20,000 to 30,000 per sex). The most frequent cancers can be reliably mapped in smaller total populations of about 5,000 to 10,000. For rarer cancers with incidence rates between 5 and 10 per 100,000 per year, a minimum total population size to generate expected numbers in the range of 10 to 16 cases in five years exceeds 100,000, or about 50,000 per sex).

Tables 2 and 3 also show the distribution of New Jersey county, municipality and census tract populations using the same ranges of population sizes. From the tables, it is



apparent that cancers with incidence rates below 10 per 100,000 can be reliably mapped at the county level in New Jersey, but that this is feasible for only a small number of municipalities. Of course, analysis of incidence data over longer time frames than five-year intervals would be more reliable. Five-year cancer data for the most common cancer type groupings (prostate, breast, lung and bronchus, and colon and rectum), may be reliably mapped for many municipalities and some census tracts in New Jersey. For the many important cancer type groupings with incidence rate between 10 and 50 per 100,000 per year, rates or relative rates may be reliably mapped at the county level, and for perhaps the 25 or 30 largest municipalities in the state.

## **Conclusion**

In New Jersey, most cancer types can be practically mapped for all counties and for a small number of municipalities. For most municipalities and some census tracts, reliable rates can only be mapped for the most common cancers.

## References

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Figure 1. Cumulative probabilities of observed number of cases in a mapped unit, for expected numbers ranging from 1 to 20.

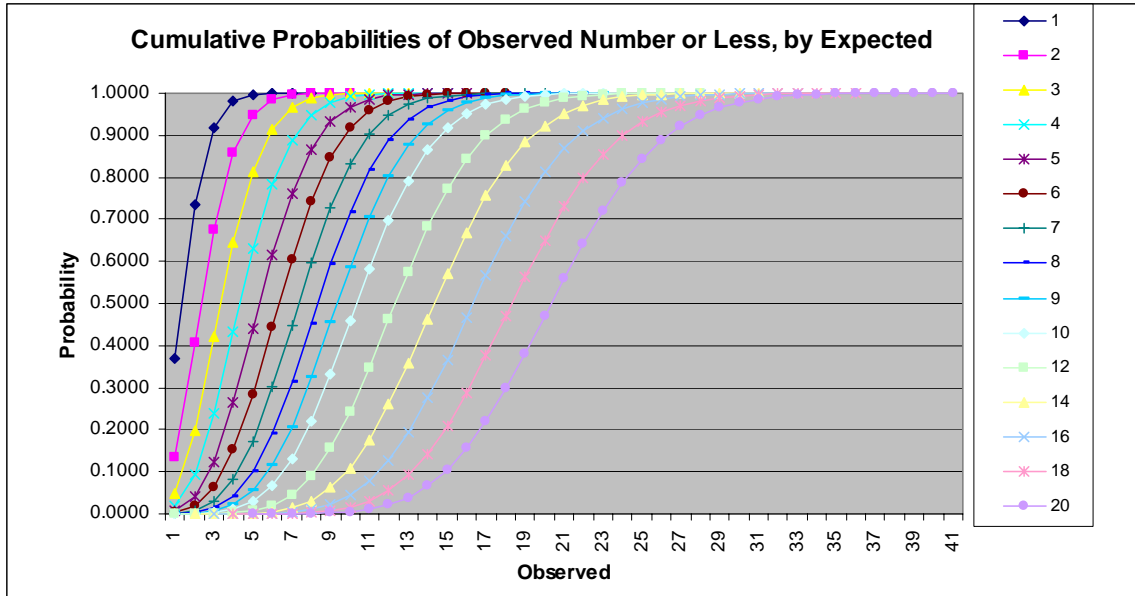


Figure 2. Range of possible rate ratios for expected numbers ranging from 1 to 20.

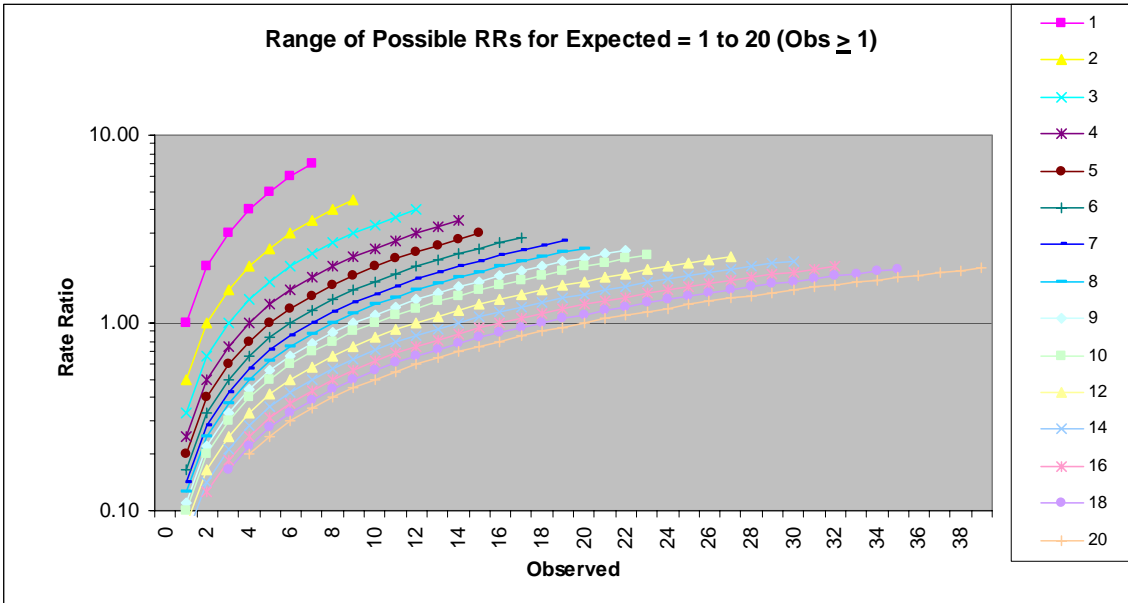


Table 1. Probability of a map unit having a rate ratio of greater than 1.5 or 2.0 by chance alone, assuming a Poisson distribution, for expected numbers ranging from 2 to 20 (only even numbers shown).

**Rate Ratio > 1.5**

<b>Expected Number</b>	<b>Observed Number for RR = 1.5</b>	<b>Observed Numbers for RR &gt; 1.5</b>	<b>Probability of Observing RR &gt; 1.5</b>
2	3	$\geq 4$	0.14
4	6	$\geq 7$	0.11
6	9	$\geq 10$	0.084
8	12	$\geq 13$	0.064
10	15	$\geq 16$	0.049
12	18	$\geq 19$	0.037
14	21	$\geq 22$	0.029
16	24	$\geq 25$	0.022
18	27	$\geq 28$	0.017
20	30	$\geq 31$	0.014

**Rate Ratio > 2.0**

<b>Expected Number</b>	<b>Observed Number for RR = 2.0</b>	<b>Observed Numbers for RR &gt; 2.0</b>	<b>Probability of Observing RR &gt; 2.0</b>
2	4	$\geq 5$	0.053
4	8	$\geq 9$	0.021
6	12	$\geq 13$	0.0088
8	16	$\geq 17$	0.0037
10	20	$\geq 21$	0.0016
12	24	$\geq 25$	0.0007
14	28	$\geq 29$	0.0003
16	32	$\geq 33$	0.0001
18	36	$\geq 37$	0.0001
20	40	$\geq 41$	< 0.0001



Table 3. Number of expected cases in males or females occurring in a 5-year observation period by cancer rate range and total population range.

Age-Standardized Annual Rate Per 100,000	Example Cancer Type By New Jersey Rate, 2000 By Sex, All Races Combined		Expected Number of Cases Based on Low Ends of Rate and Population Ranges									
			Total Population Size Range (Males plus Females)									
			Less than 1000	1000 to 2500	2500 to 5000	5000 to 10000	10000 to 25000	25000 to 50000	50000 to 100000	100000 to 250000	250000 to 500000	500000 to 1 million
	Male	Female										
<b>Greater than 100</b>	Prostate	Breast		2	6	12	25	62	125	250	625	1,250
<b>50 to 100</b>	Lung and bronchus , colon and rectum	Lung and bronchus, colon and rectum			3	6	12	31	62	125	312	625
<b>20 to 50</b>	Urinary bladder, non-Hodgkin lymphoma	Corpus uteri				2	5	12	25	50	125	250
<b>10 to 20</b>	Melanoma of the skin, kidney, leukemias, pancreas, stomach	Ovary, non-Hodgkin lymphoma, thyroid, urinary bladder, melanoma of the skin, pancreas					2	6	12	25	62	125
<b>5 to 10</b>	Esophagus, brain and other nervous system, larynx, liver, myelomas, testis, thyroid	Cervix uteri, leukemias, kidney, stomach, brain and other nervous system, myelomas						3	6	12	31	62
<b>2 to 5</b>	Soft tissue (including heart), Hodgkin lymphoma	Hodgkin lymphoma, soft tissue (including heart), esophagus, liver							2	5	12	25
<b>1 to 2</b>	Nasopharynx, bones and joints	Larynx								2	6	12
	<b># Census tracts (of 1,944)</b>		48	266	1,004	610	16					
	<b># Municipalities (of 566)</b>		27	72	93	135	150	60	25	3	1	
	<b># Counties (of 21)</b>								1	5	7	8

Appendix Table 1. Cumulative Poisson probabilities of observed number or less, given expected numbers of cases between 1 and 20. Cells highlighted in pink are cumulative probabilities for  $RR \leq 1.25$ , green are cumulative probabilities for  $RR \leq 1.5$ , and cells highlighted in yellow are cumulative probabilities for  $RR \leq 2.0$ .

OBS	Expected														
	1	2	3	4	5	6	7	8	9	10	12	14	16	18	20
0	0.3679	0.1353	0.0498	0.0183	0.0067	0.0025	0.0009	0.0003	0.0001	0.0000	0.0000				
1	0.7358	0.4060	0.1991	0.0916	0.0404	0.0174	0.0073	0.0030	0.0012	0.0005	0.0001	0.0000			
2	0.9197	0.6767	0.4232	0.2381	0.1247	0.0620	0.0296	0.0138	0.0062	0.0028	0.0005	0.0001	0.0000		
3	0.9810	0.8571	0.6472	0.4335	0.2650	0.1512	0.0818	0.0424	0.0212	0.0103	0.0023	0.0005	0.0001	0.0000	
4	0.9963	0.9473	0.8153	0.6289	0.4405	0.2851	0.1730	0.0996	0.0550	0.0293	0.0076	0.0018	0.0004	0.0001	0.0000
5	0.9994	0.9834	0.9161	0.7852	0.6160	0.4457	0.3007	0.1912	0.1157	0.0671	0.0203	0.0055	0.0014	0.0003	0.0001
6	0.9999	0.9955	0.9665	0.8894	0.7622	0.6063	0.4497	0.3134	0.2068	0.1301	0.0458	0.0142	0.0040	0.0010	0.0003
7	1.0000	0.9989	0.9881	0.9489	0.8666	0.7440	0.5987	0.4530	0.3239	0.2202	0.0895	0.0316	0.0100	0.0029	0.0008
8		0.9998	0.9962	0.9787	0.9319	0.8472	0.7291	0.5925	0.4557	0.3328	0.1550	0.0621	0.0220	0.0071	0.0021
9		1.0000	0.9989	0.9919	0.9682	0.9161	0.8305	0.7166	0.5874	0.4579	0.2424	0.1094	0.0433	0.0154	0.0050
10			0.9997	0.9972	0.9863	0.9574	0.9015	0.8159	0.7060	0.5830	0.3472	0.1757	0.0774	0.0304	0.0108
11			0.9999	0.9991	0.9945	0.9799	0.9467	0.8881	0.8030	0.6968	0.4616	0.2600	0.1270	0.0549	0.0214
12			1.0000	0.9997	0.9980	0.9912	0.9730	0.9362	0.8758	0.7916	0.5760	0.3585	0.1931	0.0917	0.0390
13				0.9999	0.9993	0.9964	0.9872	0.9658	0.9261	0.8645	0.6815	0.4644	0.2745	0.1426	0.0661
14				1.0000	0.9998	0.9986	0.9943	0.9827	0.9585	0.9165	0.7720	0.5704	0.3675	0.2081	0.1049
15					0.9999	0.9995	0.9976	0.9918	0.9780	0.9513	0.8444	0.6694	0.4667	0.2867	0.1565
16					1.0000	0.9998	0.9990	0.9963	0.9889	0.9730	0.8987	0.7559	0.5660	0.3751	0.2211
17						0.9999	0.9996	0.9984	0.9947	0.9857	0.9370	0.8272	0.6593	0.4686	0.2970
18						1.0000	0.9999	0.9993	0.9976	0.9928	0.9626	0.8826	0.7423	0.5622	0.3814
19							1.0000	0.9997	0.9989	0.9965	0.9787	0.9235	0.8122	0.6509	0.4703
20								0.9999	0.9996	0.9984	0.9884	0.9521	0.8682	0.7307	0.5591
21								1.0000	0.9998	0.9993	0.9939	0.9712	0.9108	0.7991	0.6437
22									0.9999	0.9997	0.9970	0.9833	0.9418	0.8551	0.7206
23									1.0000	0.9999	0.9985	0.9907	0.9633	0.8989	0.7875
24										1.0000	0.9993	0.9950	0.9777	0.9317	0.8432
25											0.9997	0.9974	0.9869	0.9554	0.8878
26											0.9999	0.9987	0.9925	0.9718	0.9221
27											0.9999	0.9994	0.9959	0.9827	0.9475
28											1.0000	0.9997	0.9978	0.9897	0.9657
29												0.9999	0.9989	0.9941	0.9782
30												0.9999	0.9994	0.9967	0.9865
31												1.0000	0.9997	0.9982	0.9919
32													0.9999	0.9990	0.9953
33													0.9999	0.9995	0.9973
34													1.0000	0.9998	0.9985
35														0.9999	0.9992
36														0.9999	0.9996
37														1.0000	0.9998
38															0.9999
39															0.9999
40															1.0000



Appendix Table 2. Ratios (observed/expected) for expected numbers of cases between 1 and 20.

Observed	Expected															
	1	2	3	4	5	6	7	8	9	10	12	14	16	18	20	
0																
1	1.00	0.50	0.33	0.25	0.20	0.17	0.14	0.13	0.11	0.10	0.08	0.07				
2	2.00	1.00	0.67	0.50	0.40	0.33	0.29	0.25	0.22	0.20	0.17	0.14	0.13			
3	3.00	1.50	1.00	0.75	0.60	0.50	0.43	0.38	0.33	0.30	0.25	0.21	0.19	0.17		
4	4.00	2.00	1.33	1.00	0.80	0.67	0.57	0.50	0.44	0.40	0.33	0.29	0.25	0.22	0.20	
5	5.00	2.50	1.67	1.25	1.00	0.83	0.71	0.63	0.56	0.50	0.42	0.36	0.31	0.28	0.25	
6	6.00	3.00	2.00	1.50	1.20	1.00	0.86	0.75	0.67	0.60	0.50	0.43	0.38	0.33	0.30	
7	7.00	3.50	2.33	1.75	1.40	1.17	1.00	0.88	0.78	0.70	0.58	0.50	0.44	0.39	0.35	
8		4.00	2.67	2.00	1.60	1.33	1.14	1.00	0.89	0.80	0.67	0.57	0.50	0.44	0.40	
9		4.50	3.00	2.25	1.80	1.50	1.29	1.13	1.00	0.90	0.75	0.64	0.56	0.50	0.45	
10			3.33	2.50	2.00	1.67	1.43	1.25	1.11	1.00	0.83	0.71	0.63	0.56	0.50	
11			3.67	2.75	2.20	1.83	1.57	1.38	1.22	1.10	0.92	0.79	0.69	0.61	0.55	
12			4.00	3.00	2.40	2.00	1.71	1.50	1.33	1.20	1.00	0.86	0.75	0.67	0.60	
13				3.25	2.60	2.17	1.86	1.63	1.44	1.30	1.08	0.93	0.81	0.72	0.65	
14				3.50	2.80	2.33	2.00	1.75	1.56	1.40	1.17	1.00	0.88	0.78	0.70	
15					3.00	2.50	2.14	1.88	1.67	1.50	1.25	1.07	0.94	0.83	0.75	
16						2.67	2.29	2.00	1.78	1.60	1.33	1.14	1.00	0.89	0.80	
17						2.83	2.43	2.13	1.89	1.70	1.42	1.21	1.06	0.94	0.85	
18							2.57	2.25	2.00	1.80	1.50	1.29	1.13	1.00	0.90	
19							2.71	2.38	2.11	1.90	1.58	1.36	1.19	1.06	0.95	
20								2.50	2.22	2.00	1.67	1.43	1.25	1.11	1.00	
21									2.33	2.10	1.75	1.50	1.31	1.17	1.05	
22									2.44	2.20	1.83	1.57	1.38	1.22	1.10	
23										2.30	1.92	1.64	1.44	1.28	1.15	
24											2.00	1.71	1.50	1.33	1.20	
25											2.08	1.79	1.56	1.39	1.25	
26											2.17	1.86	1.63	1.44	1.30	
27											2.25	1.93	1.69	1.50	1.35	
28												2.00	1.75	1.56	1.40	
29												2.07	1.81	1.61	1.45	
30												2.14	1.88	1.67	1.50	
31													1.94	1.72	1.55	
32													2.00	1.78	1.60	
33														1.83	1.65	
34														1.89	1.70	
35														1.94	1.75	
36															1.80	
37															1.85	
38															1.90	
39															1.95	
40															2.00	